Random processes and probability distributions

Phys 750 Lecture 16

Random processes

- Many physical processes are random in character: e.g.,
 - nuclear decay (Poisson distributed event count)

$$P(k,\tau) = \frac{e^{-\lambda\tau}(\lambda\tau)}{k!}$$

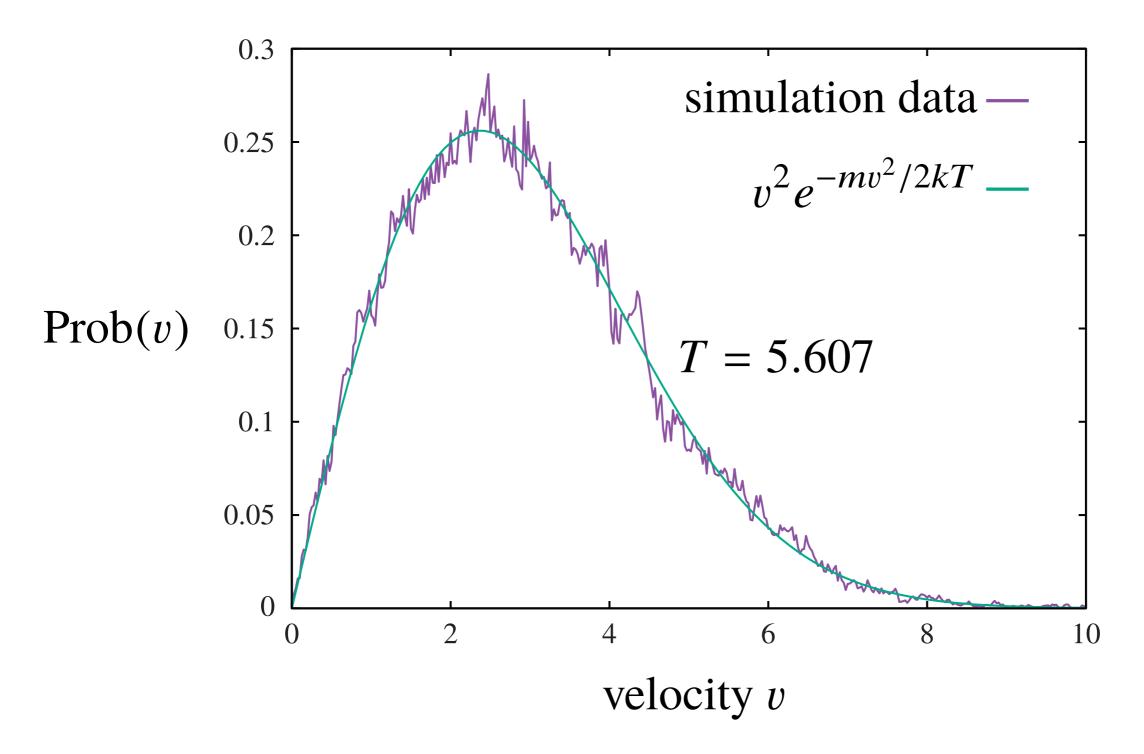
 motion of "thermalized" interacting particles (Maxwell-Boltzmann speed profile)

$$f(v) \sim v^2 \exp\left(-\frac{mv^2}{2kT}\right)$$

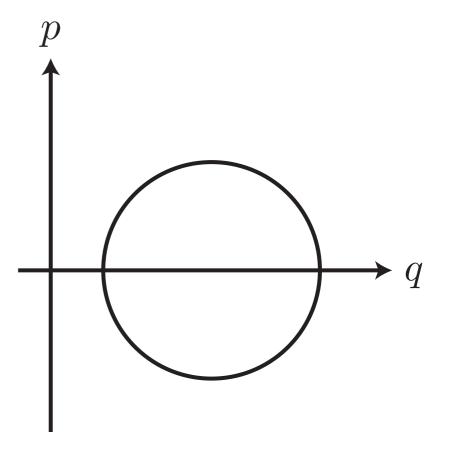
Random processes

- Probabilistic descriptions often arise from the complex behaviour of many interacting degrees of freedom
- E.g., Temperature is an emergent phenomenon:
 - it's a collective property of a large number of interacting particles
 - particles exchange energy and establish a MB distribution, characterized by a single parameter T

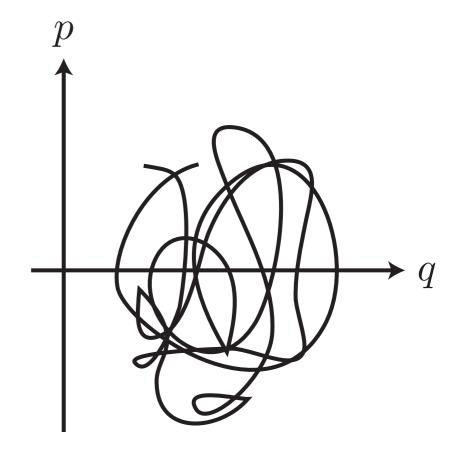
Lennard-Jones gas



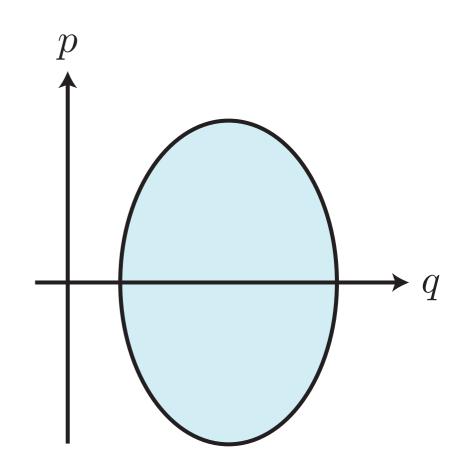
 For a large collection of noninteracting oscillators, the phase space trajectory of each individual oscillator is a closed circle



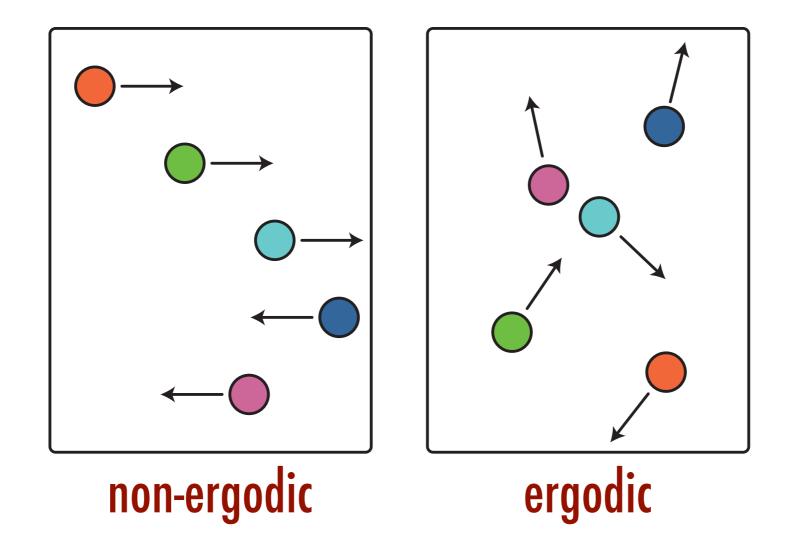
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- For a large collection of noninteracting oscillators, the phase space trajectory of each individual oscillator is a closed circle
- As we turn on interactions, the trajectories are perturbed
- The trajectory density averaged over all oscillators can be understood as a probability density



• Consider an idealized billiards system. . .



Statistical physics

- Ergodicity is the basis for statistical physics
- We assume that all states in the phase space are accessible and have equal weight
- In a system at fixed temperature, states are visited according to the Boltzmann factor:

$$Z = \sum_{n} e^{-E_n/kT}$$

Stochastic process with transition probabilities given by

$$P_{n \to m} \sim e^{-(E_n - E_m)/kT}$$

Some subtleties

- Can we prove ergodicity? Rarely
- What if the dynamics are ergodic but the timescale for traversing the phase space is slow? Glassiness
- What if the dynamics are ergodic within distinct regions of the phase space that are only tenuously connected? Rare tunnelling events
- How do we model such processes on a digital state machine that is <u>purely deterministic</u>? Pseudo-randomness

- Misleading terminology
- Is 23 a random number?
- In what sense could it be random?

- What is a random number?
 - Really no such thing
 - Loose term of art referring to a sequence of independent numbers drawn randomly from some distribution
 - Typically these are integer or real values uniformly distributed in some finite range

An infinite sequence of digits:

 $99181956211585263425870769311327827177953470784192\cdots$

- Is it random? (Humans are terrible at judging)
- Each of the digits 0-9 occurs $\frac{1}{10}$ of the time
- Each pair of two successive digits occurs $\frac{1}{100}$ of the time

- Consider the first one million digits of the sequence:
 - digit counts are distributed around the average
 - every pattern is equally probable

 $P(000000\cdots) = P(193273\cdots)$

digits are completely uncorrelated

given $000000 \cdots 0x$, $P(x = 0) = \frac{1}{10}$

- How do we generate sequences of random numbers?
- Strictly speaking, this isn't possible on a deterministic computer using finite arithmetic
- Nonetheless, it may be possible to construct long sequences with the appearance of randomness
- Probably okay if the relationship between numbers has no physical significance

- Want a random sequence of real numbers $(U_n) \in [0, 1)$
- Popular strategy:
 - use fractions $U_n = X_n/m$ built from the sequence of integers $(X_n) \in \{0, 1, 2, \dots, m-1\}$
 - Inear congruence scheme (Lehmer 1948)

$$X_{n+1} = (@X_n + c) \mod (m)$$

multiplier modulus
increment

- Recursion builds off an initial "seed" value, X_0 X_0 $X_1 = (aX_0 + c) \mod m$
 - $X_2 = (a[(aX_0 + c) \mod m] + c) \mod m$
- Requires careful choice of parameters:

$$(X_n) = (7, 6, 9, 0) (7, 6, 9, 0) \dots$$

 $X_0 = a = c = 7$
 $m = 10$
very non-random, period 4

- Any generator of the form $X_{n+1} = F(X_n)$ taking m distinct values must be periodic with period $P \le m$
- For a linear congruential generator, one can show that the period is maximum if
 - c is relatively prime to m;
 - b = a 1 is a multiple of p, for very prime p dividing m
 - b is a multiple of 4, if m is a multiple of 4.

- ${\scriptstyle \bullet}$ For the sake of efficiency, specialize to the case $\,m=2^{32}$
- Advantages:
 - each integer fits into exactly one computer word
 - the modulus (an expensive division operation) is automatically handled in hardware by overflow
- A specialized version of the maximum period rule:

c = 1 $a \equiv 5 \mod 8$

A long repeating cycle does not imply randomness:

 $(X_n) = 0, 1, 2, \dots, m-1 \quad (a = c = 1)$

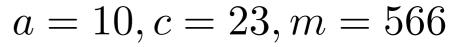
• We require weak correlation between elements,

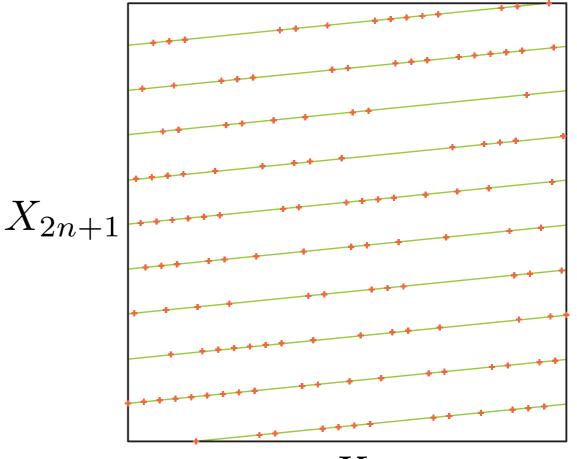
 $\langle X_j X_k \rangle \approx \langle X_j \rangle \langle X_k \rangle \ (j \neq k)_{\prime}$

especially when |j - k| is small

Judged empirically via statistical tests (Die Hard)

- Some other considerations:
 - \blacktriangleright best if a/m is not too small
 - least significant bits are more highly correlated
 - complete orbits lie in hyperplanes (Marsaglia 1968):





 X_{2n}

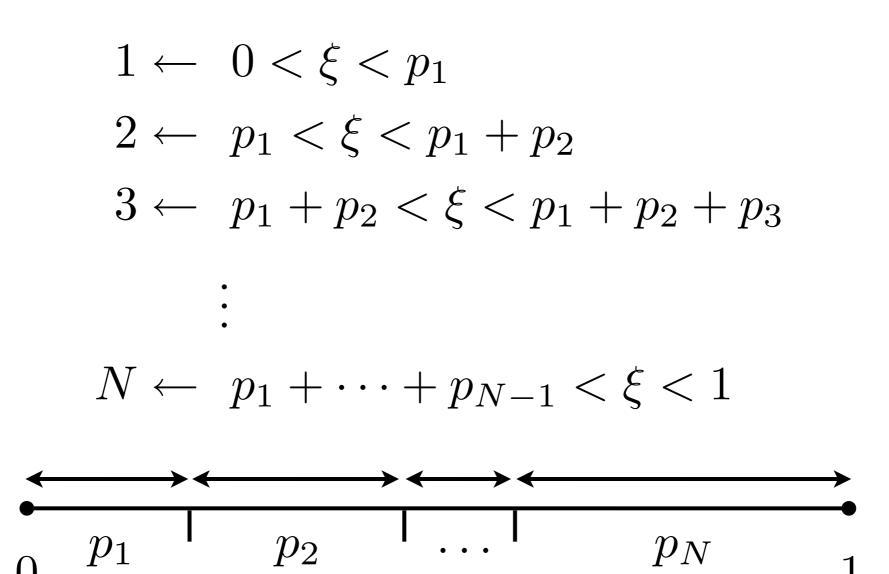
$$(X_0, X_1, \ldots, X_{q-1}), (X_q, X_{q+1}, \ldots, X_{2q-1}), \ldots$$

- Suppose there are N discrete events occur with probabilities p_1, p_2, \ldots, p_N
- Since <u>something</u> must happen, the total sum is

 $p_1 + p_2 + \dots + p_N = 1$

• Given a randomly generated number $\xi \in [0, 1]$, how can we select one of the N events?

• Each even i occupies a width p_i in the interval:



The relevant quantity is the cumulative probability,

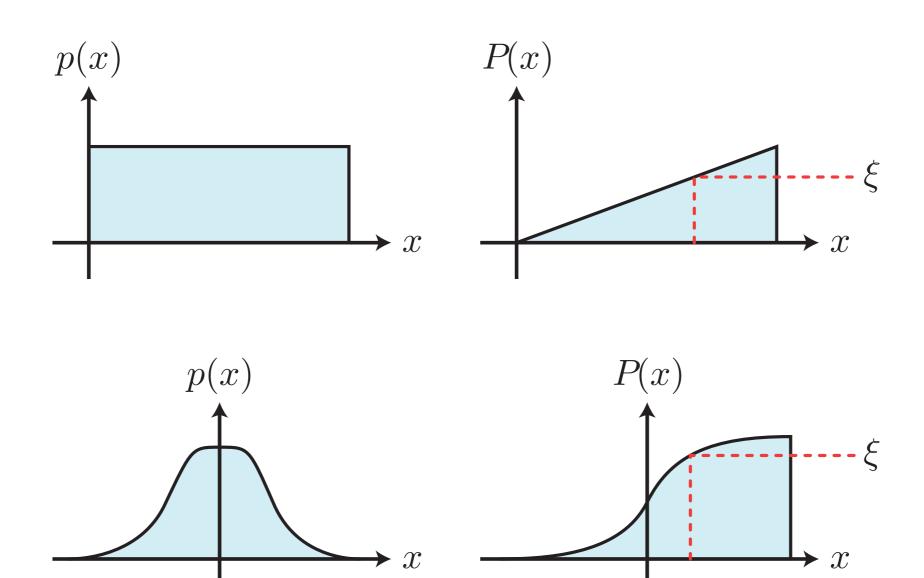
$$P_i = \sum_{j=1}^i p_j$$

 Similarly, for continuous distributions, we construct a cumulative probability distribution

$$P(x) = \int_{-\infty}^{x} dy \, p(y)$$

from the probability density p(x)

• Sampling via $x \leftarrow P^{-1}([0,1])$



Inverse transform method

• Example:
$$p(x) = \begin{cases} (1/\lambda)e^{-x/\lambda} & \text{if } 0 \le x < \infty \\ 0 & \text{if } x < 0 \end{cases}$$

• Back map gives
$$\xi \to P(x) = \int_0^x dy \, p(y) = 1 - e^{-x/\lambda}$$

• By inversion, we see that $x \leftarrow -\lambda \ln(1-\xi)$ with ξ drawn uniformly from [0,1] is equivalent to x drawn from the nonuniform distribution p(x)

Inverse transform method

- What about cases where no analytic inverse exists?
- In some cases, related multivariate distributions are invertible: e.g. Gaussian distribution

$$p(x) = \frac{e^{-x^2/2\sigma^2}}{\sqrt{2\pi\sigma^2}}$$

Consider the product

$$p(x,y) = p(x)p(y) = \frac{e^{-(x^2+y^2)/2\sigma^2}}{2\pi\sigma^2}$$

Inverse transform method

• From radial coordinates $r = \sqrt{x^2 + y^2}$ $\theta = \tan^{-1}(y/x)$

we can further transform
$$\rho = r^2/2$$
 so that
 $p(\rho, \theta) d\rho d\theta = \frac{1}{2\pi} e^{-\rho} d\rho d\theta$
 $x = \sqrt{2\rho} \cos \theta$
 $y = \sqrt{2\rho} \sin \theta$

Sampling is now possible with two random variables:

$$x \leftarrow \sqrt{-2\ln(1-\xi_1)}\cos(2\pi\xi_2)$$
 (Box-Muller)

Rejection method

- When P(x) is not easily invertible and no other tricks can be applied, try the following <u>rejection method</u>:
 - Generate a sequence (x_1, x_2, x_3, \cdots) with the elements drawn uniformly from $[x_{\min}, x_{\max}]$
 - Generate a sequence $(\xi_1, \xi_2, \xi_3, \cdots)$ with the elements drawn uniformly from $[0, p_{\max}]$
 - Discard elements from the first sequence if $p(x_i) < \xi_i$

Rejection method

- Works by throwing away results that are rare
- Some limitations: the probability distribution must be bounded and have a finite range

